# PRELIMINARY RESULTS OF STANDARDIZED CATCH RATES FOR SKIPJACK TUNA (Katsuwonus pelamis) FROM THE BRAZILIAN BAITBOAT FISHERY THROUGH 1998 

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## SUMMARY

Preliminary general linear modelling (GLM) analysis were used to attempt to standardize catch and effort data from the Brazilian baitboat fishery of skipjack in the soutwestern Atlantic. The basic data used were extracted from log-books and covered the period 1983-1998. The following main effects were incorporated into the models: fleet, year, region and quarter. Data on vessel type and vessel characteristics were also available but were not used in the analysis. It is expected that the incorporation of these addtional factors could improve the models fit.


#### Abstract

RÉSUMÉ Une analyse préliminaire par modélisation linéaire généralisée (GLM) a servi pour tenter de standardiser les données de capture et d'effort de la pêche de listao par les canneurs brésiliens dans l'Atlantique sud-ouest. Les données de base utilisées étaient extraites des carnets de pêche et couvraient la période 1983-1998. Les facteurs suivants ont été incorporés au modèle: flottille, année, région et trimestre. Des données étaient également disponibles sur les caractéristiques des bateaux, mais n'ont pas été utilisées dans les analyses. Il est escompté que l'incorporation de ces facteurs supplémentaires améliore l'ajustement du modèle.


## RESUMEN

Se aplicó el análisis preliminar del modelo lineal generalizado (GLM) en un intento de estandarizar los datos de captura y esfuerzo de la pesquería brasileña de barcos de cebo para el listado en el Atlántico sudoeste. Los datos básicos usados se obtuvieron de los cuadernos de pesca y cubrían el periodo 1983-1998. A los modelos se incorporaron los siguientes elementos: flota, año, región y trimestre. Se disponía también de datos sobre tipo y características de los barcos, si bien no se usaron en el análisis. Se confía en que la incorporación de estos factores adicionales mejore el ajuste de los modelos.

## 1. INTRODUCTION

The western skipjack stock is primarily caught by baitboats fishing off the South and Southeast coast of Brazil. Venezuelan purse seiners and baitboats also catch important portions of this stock. Minor important fisheries for skipjack are also conducted by USA purse seiners and Cuban baitboats (ICCAT, 1998).

Up to the present the SCRS has not conducted any stock assessment analysis for the W estern skipjack stock due to a lack of adequate catch and effort series of data to support such analysis.

During the ICCAT Data Preparatory Meeting for the South A tlantic Abundance Indices (ICCAT, 1995) it was tentatively decided to apply the GLM approach to develop standardized catch rates for the skipjack caught by the Brazilian baitboat fisheries. However, due to time constraints such analysis were not carried out and a recommendation was made for continued investigation in the near future.

Following this recommendation the available data from the Brazilian baitboat fishery were revised to create an appropriate data base on catch and effort covering the period from 1983 to 1998, to support the statistical analysis by the GLM procedure. The basic data consisted of information originated from mandatory submission of logbooks by fishermen of the baitboat fishery from the States of Rio de Janeiro, Santa C atarina and Rio Grande do Sul. The purpose of this paper is to present the preliminary results of the GLM analysis for standardization of skipjack catch rates, in an attempt to provide indices of abundance from the Brazilian baitboat fishery, for possible use by the SCRS to carry out stock assessment analysis for the W estern skipjack stock.

## 2. MATERIAL AND METHODS

Basic catch and effort data used in the analysis were obtained from the baitboat fishery operating off the south and southeast coast of Brazil. Skipjack is the target species for this fishery comprising about $90 \%$ of the total weight of catches. Two source of data were used: the first was comprised of logbook data which were available either as records of catch and effort data for each fishing trip, aggregated by month and one degree statistical blocks, or as individual records of daily vessel activities conducted during each fishing trip (catching live bait, moving to the fishing grounds, facing bad weather conditions, searching fish schools, days fishing); the other source of data consisted of information on the main vessel characteristics: gross registered tonnage (GRT), total length, age of the vessels and carrying capacity.

All records of catch and effort data for each vessel were matched with its corresponding vessel characteristics and the resultant data base created included the following data: fleet, vessel identity, year, month, area, days spent at sea, days spent fishing live bait, searching days, effective fishing days, catch in weight by species (skipjack, yellowfin, albacore, blackfin tuna, frigate tuna, dolphin fish and other species), vessel characteristics (total length, gross registered tonnage) and vessel building date. Nominal CPUE values were calculated as catch in weight per fishing days(including unsuccessful searching days).

As for some years the original data set of catch and effort was consisted both of aggregated and disaggregated data, in order to have a database with a uniform format to carry out more consistent analysis, all the disaggregated data were processed for the purpose of having all data in aggregated format

Two types of GLM models were used to develop standardized catch rates of skipjack in the Brazilian baitboat fishery: the $\log (C P U E+1)$ and the delta-lognormal approach of Lo et al (1992), in which the log transformed positive catch rates and the proportion of observations for which there is a positive catch are modelled separately to produce an estimated abundance index. The analysis were conducted following the methodology applied at the ICCAT Data Preparatory M eeting for the South A tlantic A bundance Indices (ICCAT, 1995). Further details of this method are presented by Cramer and Scott (1993, 1997).

In order to allow the application of the GLM model based on the delta-lognormal assumption, each aggregated data record with both fishing days with positive catch and fishing days with zero catch were aggregated separately by all combination of categorical variables which were assumed to have influence on skipjack catch rates. Following this procedure, all observations in the original data set were aggregated by fleet, vessel identity, year, month, quarter, region, $1^{\circ} \times 1^{\circ}$ area, GRT and vessel length classes.

For this fishery it was assumed that the main factors which influence CPUE are: year, season, area and vessel characteristics. Other factors with considerable influence on catch rates are the bait used, the type of schools from which the catches were made and the fishermen's skill (ICCAT, 1995). F or the present analysis the main effects used in the GLM model were: year, fleet, quarter and area. The zero catch records were included in the model and observations with extremely small effort were excluded from the analysis.

## 3. RESULTS AND DISCUSSION

The original data set available for the analysis consisted of 12,455 observations. The nominal data were planned to be examined by plotting CPUE versus all the available variable which could influence CPUE to determine which variable should be considered as a factor in the model.. However, it was not possible to investigate the effect of all the possible variables on skipjack catch rates and it was decided to restrict the analysis to the following main effects: fleet, year, quarter and region. Tables 1 and 2 presents summaries of these observations by fleet, year, region and quarter, considering all two way combinations of these main effects.

Figures 1 through 6 show plots of CPUE versus year and area and CPUE versus year and season. Two geographical fishing areas were defined based on nominal catch distributions: north area, comprising all one degree blocks to the N orth of latitude $28^{\circ} \mathrm{S}$, and south area comprising all the blocks situated to the south of this latitudinal line. The main factors considered in defining these two regions were:(1) consistently higher skipjack CPUE in the south area than in the north area, as can be seen from figures 1 through 3; and (2) higher proportion of catches of yellowfin tuna in the North area in comparison with the south area. There are also indications that in the north area some baitboats use to operate around several oil rig platforms placed in the area. It is supposed that the higher proportion of yellowfin catches in this area results from the higher aggregation of yellowfin schools around the platforms. Fishing on this tuna aggregations is advantageous for the baitboat fishermen as they avoid spending time searching for tuna schools in the open sea.

The examination of plots of nominal CPUE for each fleet, by year and season, showed a consistent pattern of higher CPUE during the first quarter, for both leased fleet and Santa C atarina fleet, and smaller CPUE during the third quarter for all fleets (Fig. 4 and 6). For the Rio de Janeiro fleet, CPUE trends by season are not consistent from 1983 through 1998, however it is noted that the highest CPUE's have mainly occurred in the first and second quarters (fig. 5).

Although no consistent relationships could be observed between quarterly skipjack and yellowfin catch rates for leased baitboats and for national baitboats based at Santa Catarina, due to the low yellowfin catches obtained by these fleets, skipjack catch rates varied inversely with yellowfin catch rates for Rio de J aneiro fleet (figure 7).

V essel characteristics were considered to be an important factor influencing skipjack catch rates. However, they were not included in the analysis because of some difficulties in defining a vessel characteristic strata to use as a categorical variable, in which vessels pertaining to each
fleet could be included. This is the case of the leased baitboat fleet, in which all the vessels have homogeneous characteristics for length and GRT. In this way, to avoid the exclusion of a great portion of the available data in the analysis, no vessel characteristics factor were explicitly considered. Although, it can be considered that vessel characteristics could explain one part of fishing success, as a result of having higher bait capacity and autonomy at sea, with bigger vessels being expected to show better fishing performance, fishermen experience and their ability to locate and concentrate fish schools are possibly more important factors explaining fishing success for this fishery. For the purpose of the analysis it was assumed that all these factors were implicitly taken into account when the vessels were classified into three separate fleets. As each fleet show different vessel characteristics and it is supposed that the biggest vessels will have a more experienced crew, the application of fishermen's knowledge on the environmental factors which influence the distribution and concentration of fish schools, would allow them to choose the best area and season for fishing.

The better fishing performance of the leased fleet, comprised of the biggest vessels (over 150 GRT), in comparison with the national baitboats has been shown by IBAMA(1996) who reported that the mean annual skipjack catch rates for the leased vessels were consistently higher than for the national vessels, when fishing at the same areas and seasons.

Plots of nominal catch rates of skipjack by year and fleet showed a similar pattern of fluctuation in annual CPUE among fleets (fig. 8). For the leased fleet skipjack CPUE was consistently higher than CPUE for the other fleets. For the Brazilian fleet the examination of plots of nominal data also shown that CPUE for Santa Catarina fleet was always higher than CPUE for Rio de Janeiro fleet. It is also suggested that differences in skipjack CPUE between Rio de Janeiro and Santa C atarina fleets have increased since 1989.

Tables 3 and 4 present summaries of analysis of variance for each GLM model analysis as well as plots of histograms of the standardized residuals and the estimated skipjack abundance index with associated statistics. The plot of the fit residuals of both GLM analysis were approximately normally distributed, which suggest that both models fit could be considered acceptable. However, a poor fit (in terms of $\mathrm{r}^{2}$ ) was presented by the lognormal model, and it is possible that improvements in $r^{2}$ would be expected by adding factors for vessel type and/or vessel characteristics and also for interaction terms.

The least square adjusted CPUE with its corresponding confidence limits are plotted in figures 9 and 10 and their estimated values are show in table 3 and 4. Although no consistent trends could be noted from standardized CPUE estimated by the lognormal model, when plotting the standardized CPUE estimated by the delta-lognormal model an apparent declining trend is observed from 1985 through 1995 and an increase is noted from 1996 through 1998.

Thanks are extended to Dr. Steven Turner for providing the SAS code used to develop the GLM analysis and for introducing us the basic principles of GLM analysis for development of standardized catch rates. .

## REFERENCE CITED

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Table 1. Number of observations by fleet, year, quarter and region used in the GLM analysis by the Lognormal model

TABLE OF YEAR BY REGION
( 1= Leased baitboat fleet; 2= Rio Janeiro fleet; 3= Santa Catarina fleet)

| YEAR FLEET |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Frequency | 1 | 2 | 3 | Total |
| 83 | 198 | 585 | 62 | 845 |
| 84 | 245 | 421 | 192 | 858 |
| 85 | 290 | 479 | 184 | 953 |
| 86 | 251 | 346 | 74 | 671 |
| 87 | 259 | 205 | 52 | 518 |
| 88 | 241 | 76 | 39 | 358 |
| 89 | 214 | 83 | 164 | 461 |
| 90 | 182 | 193 | 85 | 460 |
| 91 | 157 | 107 | 147 | 411 |
| 92 | 208 | 85 | 216 | 509 |
| 93 | 201 | 107 | 301 | 612 |
| 94 | 179 | 102 | 243 | 524 |
| 95 | 290 | 168 | 208 | 666 |
| 96 | 194 | 120 | 153 | 467 |
| 97 | 152 | 104 | 411 | 667 |
| 98 | 164 | 90 | 388 | 642 |
| Total | 3425 | 3271 | 2922 | 9618 |

TABLE OF YEAR BY QUATR
( $1=$ Lesead baitboat fleet; 2= Rio Janeiro fleet; 3= Santa Catarina)

| FLEET |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Frequency | 1 | 2 | 3 | 4 | Total |
| 1 | 1008 | 1169 | 430 | 818 | 3425 |
| 2 | 866 | 980 | 693 | 752 | 3271 |
| 3 | 1139 | 982 | 389 | 412 | 2922 |
| Total | 3013 | 3151 | 1512 | 1962 | 9518 |

TABLE OF YEAR BY QUATR

| YEAR |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Frequency | QUATR |  |  |  |  |
| 83 | 235 | 223 | 166 | 221 | 845 |
| 84 | 248 | 236 | 171 | 203 | 858 |
| 85 | 258 | 305 | 169 | 221 | 953 |
| 86 | 196 | 212 | 122 | 141 | 671 |
| 87 | 110 | 168 | 106 | 132 | 518 |
| 88 | 113 | 136 | 46 | 59 | 358 |
| 89 | 129 | 151 | 63 | 118 | 461 |
| 90 | 120 | 142 | 93 | 105 | 460 |
| 91 | 168 | 130 | 35 | 78 | 411 |
| 92 | 153 | 214 | 57 | 85 | 509 |
| 93 | 233 | 173 | 99 | 107 | 612 |
| 94 | 183 | 164 | 54 | 98 | 524 |
| 95 | 216 | 206 | 108 | 136 | 666 |
| 96 | 188 | 149 | 80 | 70 | 467 |
| 97 | 252 | 245 | 76 | 94 | 667 |
| 98 | 211 | 247 | 85 | 99 | 642 |
| Total | 3031 | 3131 | 1512 | 1962 | 9618 |

TABLE OF YEAR BY REGION
FLEET ( 1 = Leasea baitboat fleet; 2= Rio Janeiro fleet; 3= Santa Catarina fleet)
REGION ( $1=$ North of $28^{\circ} \mathrm{S} ; 2=$ South $28^{\circ} \mathrm{S}$ )
FLEET REGION

| Frequency |  | 1 | 2 |
| :--- | :--- | :--- | :--- |
| Total |  |  |  |
| 1 | 1572 | 1853 | 3425 |
| 2 | 3245 | 26 | 3271 |
| 3 | 1959 | 963 | 2922 |
| Total | 6776 | 2482 | 9618 |

TABLE OF YEAR BY REGION
( $1=$ North of $28^{\circ} \mathrm{S} ; 2=$ South of $28^{\circ} \mathrm{S}$ )
YEAR REGION

| Frequency | 1 | 2 | Total |
| ---: | ---: | ---: | ---: |
| 83 | 766 | 79 | 845 |
| 84 | 654 | 204 | 858 |
| 85 | 695 | 257 | 953 |
| 86 | 497 | 174 | 671 |
| 87 | 333 | 163 | 516 |
| 88 | 228 | 128 | 356 |
| 89 | 295 | 166 | 461 |
| 90 | 338 | 122 | 460 |
| 91 | 237 | 174 | 411 |
| 92 | 320 | 189 | 509 |
| 93 | 365 | 247 | 612 |
| 94 | 323 | 201 | 524 |
| 95 | 427 | 239 | 666 |
| 96 | 340 | 127 | 467 |
| 97 | 511 | 156 | 667 |
| 98 | 446 | 196 | 642 |
| 98 | 6776 | 2842 | 9618 |
| Total |  |  |  |

TABLE OF YEAR BY REGION
QUATR REGION

| Frequency | 1 | 2 |  |
| ---: | ---: | ---: | ---: |
| 1 | 1569 | 1444 | 3013 |
| 2 | 2441 | 690 | 3131 |
| 3 | 1397 | 115 | 1512 |
| Total | 4 | 1369 | 593 |

Table 2. Number of observations by fleet, year, quarter and region used in the GLM analysis by the Lognormal model

TABLE OF YEAR BY REGION
FLEET( 1= Leased baitboat fleet; 2= Rio Janeiro fleet; 3= Santa Catarina Fleet).

| YEAR |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
| FLEET |  |  |  |  |
| Frequency | 1 | 2 | 3 | Total |
| 83 | 242 | 877 | 78 | 1197 |
| 84 | 310 | 592 | 255 | 1157 |
| 85 | 366 | 679 | 265 | 1310 |
| 86 | 316 | 497 | 112 | 639 |
| 87 | 310 | 250 | 69 | 659 |
| 88 | 297 | 98 | 61 | 448 |
| 89 | 256 | 99 | 221 | 576 |
| 90 | 215 | 211 | 98 | 522 |
| 91 | 189 | 120 | 198 | 507 |
| 92 | 253 | 96 | 294 | 540 |
| 93 | 255 | 150 | 369 | 794 |
| 94 | 204 | 155 | 325 | 694 |
| 95 | 392 | 223 | 266 | 881 |
| 96 | 226 | 125 | 182 | 534 |
| 97 | 180 | 115 | 523 | 819 |
| 98 | 195 | 104 | 515 | 814 |
| Total | 4206 | 4410 | 3839 | 12455 |

TABLE OF YEAR BY QUATR
FLEET ( 1 = Lesead baitboat fleet; 2= Rio Janeiro fleet; 3= Santa Catarina fleet).

| FLEET |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Frequency | QUATR |  |  |  |  |
| 1 | 1197 | 1482 | 554 | 983 | 4206 |
| 2 | 1176 | 1358 | 930 | 946 | 4410 |
| 3 | 1466 | 1319 | 525 | 529 | 3839 |
| Total | 3829 | 4159 | 2009 | 2458 | 12455 |

TABLE OF YEAR BY QUATR

| YEAR QUATR |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Frequency | 1 | 2 | 3 | 4 | Total |
| 83 | 324 | 335 | 246 | 291 | 1197 |
| 84 | 335 | 325 | 235 | 261 | 1157 |
| 85 | 341 | 440 | 229 | 300 | 1310 |
| 86 | 268 | 309 | 168 | 180 | 639 |
| 87 | 136 | 207 | 188 | 158 | 659 |
| 88 | 143 | 105 | 62 | 76 | 448 |
| 89 | 153 | 191 | 85 | 147 | 576 |
| 90 | 133 | 165 | 99 | 125 | 522 |
| 91 | 199 | 168 | 46 | 94 | 507 |
| 92 | 198 | 272 | 71 | 99 | 540 |
| 93 | 301 | 235 | 126 | 132 | 794 |
| 94 | 324 | 262 | 79 | 119 | 694 |
| 95 | 283 | 286 | 149 | 163 | 881 |
| 96 | 212 | 178 | 67 | 77 | 534 |
| 97 | 313 | 294 | 97 | 115 | 819 |
| 98 | 255 | 326 | 112 | 121 | 814 |
| Total | 3829 | 4159 | 2009 | 2458 | 12455 |

## TABLE OF YEAR BY REGION

FLEET ( 1 = Lesead baitboat fleet; 2= Rio Janeiro fleet; 3= Santa Catarina fleet).
REGION ( $1=$ North of $28^{\circ} \mathrm{S} ; 2=$ South $28^{\circ} \mathrm{S}$ )
FLEET REGION

| Frequency | 1 | 2 | Total |
| :--- | :--- | :--- | :--- |
| 1 | 1928 | 2276 | 4206 |
| 2 | 4372 | 38 | 4410 |
| 3 | 2553 | 1285 | 3839 |
| Total | 8853 | 3602 | 12455 |

TABLE OF YEAR BY REGION
( $1=$ North of $28^{\circ} \mathrm{S} ; 2=$ South of $28^{\circ} \mathrm{S}$ )
YEAR REGION

| Frequency | 1 | 2 | Total |
| ---: | ---: | ---: | ---: |
| 83 | 1102 | 95 | 1197 |
| 84 | 884 | 273 | 1157 |
| 85 | 967 | 343 | 1310 |
| 86 | 687 | 236 | 639 |
| 87 | 407 | 232 | 659 |
| 88 | 286 | 160 | 448 |
| 89 | 374 | 202 | 576 |
| 90 | 382 | 140 | 522 |
| 91 | 288 | 219 | 507 |
| 92 | 394 | 246 | 540 |
| 93 | 472 | 322 | 794 |
| 94 | 454 | 240 | 694 |
| 95 | 570 | 311 | 881 |
| 96 | 382 | 152 | 534 |
| 97 | 632 | 187 | 819 |
| 98 | 572 | 242 | 814 |
| 98 | 8853 | 3602 | 12455 |
| Total |  |  |  |

TABLE OF YEAR BY REGION
REGION: $1=$ North of $28^{\circ} \mathrm{S}$;
$2=$ South of $28^{\circ} \mathrm{S}$;
QUATR REGION

| Frequency | 1 | 2 | Total |
| ---: | ---: | ---: | ---: |
| 1 | 2049 | 1780 | 3829 |
| 2 | 3229 | 930 | 4159 |
| 3 | 1848 | 161 | 2009 |
| 4 | 1727 | 731 | 2458 |
| Total | 8853 | 3602 | 12455 |

Table 3. Analytical results from the GLM analysis (delta-log normal model) of skipjack catch rates in the Brazilian baitboat fishery.

## G.Mon proportion positi ves

General Li near Mbdel s Procedure
C ass Level Infornati on


Nunber of observations in data set $=320$

Dependent Vari abl e: POS

|  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Source | DF | Sum of <br> Squares | Mean <br> Square | F Val ue | Pr > F |
| Model | 23 | 1.1962570 | 0.0520112 | 4.65 | 0.0001 |
| Error | 296 | 3.3114268 | 0.0111873 |  |  |
|  |  |  |  |  |  |
| Corrected Total | 319 | 4.5076838 |  |  |  |
|  |  |  | C. V. | Root MSE | POS Mean |
|  | R-Square | 23.03619 | 0.1058 | 0.4591 |  |


| Source | DF | Type III SS | Mean Square | F Val ue | Pr $>$ F |
| :--- | ---: | ---: | ---: | ---: | ---: |
| YEAR | 15 | 0.4140086 | 0.0276006 | 2.47 | 0.0020 |
| FLEET | 2 | 0.0843129 | 0.0421565 | 3.77 | 0.0242 |
| QAATR | 3 | 0.1970041 | 0.0656680 | 5.87 | 0.0007 |
| REG ON | 1 | 0.0676939 | 0.0676939 | 6.05 | 0.0145 |
| FLEET*REG ON | 2 | 0.1901039 | 0.0950519 | 8.50 | 0.0003 |

STANDARD ZED RESI UALS

G.Mon positi ve catches

Gener al Li near Mbdel s Procedure
Cass Level Infornati on

| C ass | Level s | Val ues |
| :--- | ---: | :--- |
| YEAR | 16 | 83848586878889909192939495969798 |
| FLEET | 3 | 1 |
| QAATR | 4 | 1 |


| Dependent Vari abl e: LSKJ ©R |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Sum of | Mean |  |  |
| Source | DF | Squares | Square | F Val ue | Pr $>\mathrm{F}$ |
| Mbdel | 21 | 3096. 2119 | 147. 4387 | 99.50 | 0. 0001 |
| Error | 7500 | 11113. 4198 | 1. 4818 |  |  |
| Corrected Total | 7521 | 14209. 6317 |  |  |  |
|  | R-Square | C. V. | Root MSE | LSKJ CR Mean |  |
|  | 0. 217895 | 86. 51012 | 1. 2173 |  | 1. 4071 |
| Source | DF | Type III SS | Mean Square | F Val ue | Pr $>\mathrm{F}$ |
| YEAR | 15 | 413. 1000 | 27. 5400 | 18. 59 | 0. 0001 |
| FLEET | 2 | 1261. 4910 | 630.7455 | 425. 66 | 0. 0001 |
| QATR | 3 | 391. 0955 | 130. 3652 | 87. 98 | 0. 0001 |
| REG ON | 1 | 30. 3135 | 30. 3135 | 20. 46 | 0. 0001 |

st andar di zed resi dual s

| Vari abl e=SRESI D |  |  |  |
| :---: | :---: | :---: | :---: |
|  | H stogram |  | \# Boxpl ot |
| 2. 75 |  |  |  |


| ****** | 225 | \| |
| :---: | :---: | :---: |
| ******************* | 703 | \| |
| *************************************** | 1469 | +----+ |
| ************************************************ | 1789 | *-----* |
| ************************************** | 1414 | +-+-+ |
| ********************** | 800 | \| |
| ************* | 459 | \| |
| ******** | 291 | \| |
| **** | 151 | 0 |
| *** | 87 | 0 |
| ** | 47 | 0 |
| * | 21 | 0 |
| * | 10 | * |
| * | 2 | * |
| * | 3 | * |
| -5. $75+$ | 1 | * |

cpue is uncorrected nodel cpue fromgl mon positive catches. ppos is the nodel estimated proportion positive. I NDEX is the annual, standardi zed CPUE via the Lo method with a standard error of SE_I and CV of CV_I.

| YEAR | CPUE | PPOS | I INEX | SE_I | CV_I | L80\% | U80\% |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |  |  |
| 83 | 5.12639 | 0.55968 | 6.04474 | 1.43663 | 0.23767 | 4.49891 | 8.12173 |
| 84 | 2.81340 | 0.44861 | 2.68421 | 0.23189 | 0.08639 | 2.40785 | 2.99229 |
| 85 | 5.52343 | 0.58639 | 6.87439 | 0.51511 | 0.07493 | 6.25588 | 7.55405 |
| 86 | 6.03268 | 0.50732 | 6.49833 | 0.55924 | 0.08606 | 5.83172 | 7.24114 |
| 87 | 4.69033 | 0.59044 | 5.86612 | 0.51658 | 0.08806 | 5.25117 | 6.55309 |
| 88 | 4.09586 | 0.56664 | 4.90482 | 0.51904 | 0.10582 | 4.29415 | 5.60233 |
| 89 | 4.32227 | 0.58191 | 5.32516 | 0.49342 | 0.09266 | 4.73955 | 5.98313 |
| 90 | 3.59857 | 0.63846 | 4.85972 | 0.44167 | 0.09088 | 4.33491 | 5.44806 |
| 91 | 3.17426 | 0.63010 | 4.23053 | 0.39593 | 0.09359 | 3.76091 | 4.75878 |
| 92 | 3.30523 | 0.61743 | 4.31887 | 0.39452 | 0.09135 | 3.85023 | 4.84454 |
| 93 | 3.01721 | 0.61439 | 3.92745 | 0.33291 | 0.08477 | 3.53028 | 4.36930 |
| 94 | 3.52553 | 0.57476 | 4.29456 | 0.37800 | 0.08802 | 3.84456 | 4.79723 |
| 95 | 3.00152 | 0.56967 | 3.62530 | 0.31222 | 0.08612 | 3.25315 | 4.04002 |
| 96 | 3.62274 | 0.64467 | 4.93949 | 0.45683 | 0.09249 | 4.39724 | 5.54860 |
| 97 | 4.52210 | 0.60594 | 5.80700 | 0.48963 | 0.08432 | 5.22269 | 6.45668 |
| 98 | 3.86355 | 0.69715 | 5.70163 | 0.46518 | 0.08159 | 5.14550 | 6.31786 |

I ndex with $80 \% \mathrm{C}$

Table 4. Analytical results from the GLM analysis (delta-log normal model) of skipjack catch rates in the Brazilian baitboat fishery.

## G.Mon proportion positi ves

General Li near Model s Procedure
C ass Level I nf or mati on

| C ass | Level s | Val ues |
| :---: | :---: | :---: |
| YEAR | 16 | 838485 |
| FLEET | 3 | 123 |
| QUATR | 4 | 1234 |
| REG ON | 2 | 12 |

Nunber of observations in data set $=320$
Dependent Vari abl e: POS

|  |  | Sumof | Mean |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Source | DF | Squares | Square | F Val ue | Pr $>\mathrm{F}$ |
| Mbdel | 23 | 1. 1962570 | 0. 0520112 | 4. 65 | 0. 0001 |
| Error | 296 | 3. 3114268 | 0. 0111873 |  |  |
| Corrected Total | 319 | 4. 5076838 |  |  |  |
|  | R-Square | C. V. | Root MSE |  | POS Mean |
|  | 0. 265382 | 23. 03619 | 0. 1058 |  | 0. 4591 |
| Source | DF | Type III SS | Mean Square | F Val ue | Pr $>\mathrm{F}$ |
| YEAR | 15 | 0. 4140086 | 0. 0276006 | 2. 47 | 0. 0020 |
| FLEET | 2 | 0. 0843129 | 0. 0421565 | 3. 77 | 0. 0242 |
| QATR | 3 | 0. 1970041 | 0. 0656680 | 5. 87 | 0. 0007 |
| REG ONFLEET*REG ON | 1 | 0. 0676939 | 0. 0676939 | 6. 05 | 0. 0145 |
|  | 2 | 0. 1901039 | 0. 0950519 | 8. 50 | 0. 0003 |

STANDARD ZFD RESI UALS


GMon posi ti ve catches
General Li near Mbdel s Procedure
C ass Level Inf or nati on

| C ass | Level s | Val ues |
| :---: | :---: | :---: |
| YEAR | 16 | 838485 |
| FLEET | 3 | 123 |
| QUTR | 4 | 1234 |
| REG ON | 2 | 12 |

Nunber of observations in data set $=7522$
Dependent Vari abl e: LSKJ CR

|  |  | Sumof | Mean |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Source | DF | Squares | Square | F Val ue | Pr $>$ F |
| Model | 21 | 3096.2119 | 147.4387 | 99.50 | 0.0001 |
| Error | 7500 | 11113.4198 | 1.4818 |  |  |
|  |  |  |  |  |  |
| Corrected Tot al | 7521 | 14209.6317 |  |  |  |
|  |  |  | C. V. | Root MSE | LSK CR Mean |
|  | R-Square | 86.51012 | 1.2173 |  | 1.4071 |
|  | 0.217895 |  |  |  |  |
| Source | DF | Type III SS | Mean Square | F Val ue | Pr > F |


| YEAR | 15 | 413.1000 | 27.5400 | 18.59 | 0.0001 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| FLEET | 2 | 1261.4910 | 630.7455 | 425.66 | 0.0001 |
| QATR | 3 | 391.0955 | 130.3652 | 87.98 | 0.0001 |
| REG QN | 1 | 30.3135 | 30.3135 | 20.46 | 0.0001 |

st andar di zed resi dual s

| Vari abl e=SRESI D |  |  |  |
| :---: | :---: | :---: | :---: |
| H stogram |  | \# | Boxpl ot |
| 2. $75+^{*}$ | 5 | 0 |  |
| ** | 45 | 0 |  |
| ****** | 225 | , |  |
| ******************* | 703 | \| |  |
| *************************************** | 1469 | +----+ |  |
| ************************************************ | 1789 | *-----* |  |
| ***************** | 1414 | +-+-+ |  |
| ********************** | 800 | \| |  |
| ************* | 459 | \| |  |
| ******** | 291 | \| |  |
| **** | 151 | 0 |  |
| *** | 87 | 0 |  |
| ** | 47 | 0 |  |
| * | 21 | 0 |  |
| * | 10 | * |  |
| * | 2 | * |  |
| * | 3 | * |  |
| -5. 75+* | 1 | * |  |

cpue is uncorrected nodel cpue fromgl mon positive catches. ppos is the nodel estimated proportion positive.INDEX is the annual, standardi zed CPUE via the Lo nethod with a standard error of SE_I and CV of CV_I.

| YEAR | CPUE | PPOS | I NDEX | SE_I | CV_I | L80\% | U80\% |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 83 | 5. 12639 | 0. 55968 | 6. 04474 | 1. 43663 | 0. 23767 | 4. 49891 | 8. 12173 |
| 84 | 2. 81340 | 0. 44861 | 2. 68421 | 0. 23189 | 0. 08639 | 2. 40785 | 2. 99229 |
| 85 | 5. 52343 | 0. 58639 | 6. 87439 | 0. 51511 | 0. 07493 | 6. 25588 | 7. 55405 |
| 86 | 6. 03268 | 0. 50732 | 6. 49833 | 0. 55924 | 0. 08606 | 5. 83172 | 7. 24114 |
| 87 | 4. 69033 | 0. 59044 | 5. 86612 | 0. 51658 | 0. 08806 | 5. 25117 | 6. 55309 |
| 88 | 4. 09586 | 0. 56664 | 4. 90482 | 0. 51904 | 0. 10582 | 4. 29415 | 5. 60233 |
| 89 | 4. 32227 | 0. 58191 | 5. 32516 | 0. 49342 | 0. 09266 | 4. 73955 | 5. 98313 |
| 90 | 3. 59857 | 0. 63846 | 4. 85972 | 0. 44167 | 0. 09088 | 4. 33491 | 5. 44806 |
| 91 | 3. 17426 | 0. 63010 | 4. 23053 | 0. 39593 | 0. 09359 | 3. 76091 | 4. 75878 |
| 92 | 3. 30523 | 0. 61743 | 4. 31887 | 0. 39452 | 0. 09135 | 3. 85023 | 4. 84454 |
| 93 | 3. 01721 | 0. 61439 | 3. 92745 | 0. 33291 | 0. 08477 | 3. 53028 | 4. 36930 |
| 94 | 3. 52553 | 0. 57476 | 4. 29456 | 0. 37800 | 0. 08802 | 3. 84456 | 4. 79723 |
| 95 | 3. 00152 | 0. 56967 | 3. 62530 | 0. 31222 | 0. 08612 | 3. 25315 | 4. 04002 |
| 96 | 3. 62274 | 0. 64467 | 4. 93949 | 0. 45683 | 0. 09249 | 4. 39724 | 5. 54860 |
| 97 | 4. 52210 | 0. 60594 | 5. 80700 | 0. 48963 | 0. 08432 | 5. 22269 | 6. 45668 |
| 98 | 3. 86355 | 0. 69715 | 5. 70163 | 0. 46518 | 0. 08159 | 5. 14550 | 6. 31786 |



Fig. 1 - Skipjack catch rates by year and area, Japanese leased baitboat fleet


Fig. 2 - Skipjack catch rates by year and area, Rio de Janeiro-based baitboats


Fig. 3 - Skipjack catch rates by year and area, Santa Catarina-based baitboats


Fig. 4 - Skipjack catch rates by year and quarter, leased baitboats


Fig. 5 - Skipjack catch rates by year and quarter, baitboat fleet based at Rio de Janeiro


Fig. 6 - Skipjack catch rates by year and quarters, baitboat fleet based at Santa Catarina

Fig. 7 Catch rates of skipjack and yellowfin by quarters from baitboats based at Rio de Janeiro, during the period 1983-98


Fig. 8. Nominal CPUE trends for skipjack, by each baitboat fleet, during the period 19831998.


Fig. 9 Annual change of skipjack standardized CPUE and estimated 80\% confidence intervals, by GLM (delta-log normal) in the Brazilian baitboat fishery


Fig. 10 Annual Change of standardized CPUE and estimated 95\% confidence intervals, by GLM (log normal), for skipjack in the Brazilian baitboat fishery


